Parallelization for Markov chains Monte Carlo with heterogeneous runtimes



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Contributions

- Even with parallelization, adding more chains can reduce the efficiency of MCMC, because we always wait for the slowest chain to finish.
- · We propose instead to only wait for the fastest chains.

Waiting for the slowest chains

Case study: ODE-based models

- Models in pharmacometrics integrate a scientific model inside a statistical model.
- The likelihood is parametrized by the solution to an ODE, motivated by a biological model [1].
- In this poster, we will focus on the Michaelis-Menten Model [2].
- Depending on where we are in the parameter space (Figure 1.) Solving ODEs can be numerically challenging and take an excessive amount of time.
- This results in chains with wildly heterogeneous runtimes (Figure 2.), and a diminished return in efficiency when running many chains and waiting for the slowest chain to finish (Figure 3.).

Details of the experiment

- We fit the model using Stan's Hamiltonian Monte Carlo, with 500 warmup and 500 sampling iterations. Convergence of the chains is evaluated using \hat{R} .
- The ODE is solved using a Runge-Kutta 4th/5th order integrator (RK 45). Using the more conservative backward differentiation (BDF) integrator leads to a slower average runtime but more stable behavior. RK45 is usually more efficient to solve non-stiff ODEs whereas BDF is faster for stiff ones. A stiff ODE is unstable numerically when a small perturbation is applied to its parameters.
- In this model we use **weakly informative priors** and **over-dispersed initializations.** Stronger priors (if available) and better initialization can stabilize the chains.

Observation The probability of having a « lagging » chain increases with the total number of chains.



Figure 1. [2] Evolution of 2 chains (orange and black dots) across the parameter space.
The ODE behavior varies across the space, and it can be stiff or non-stiff depending on its location. The blue ellipsoid represents the sampling region, where the posterior probability mass concentrates.



Figure 2. Runtime of 30 Markov chains when fitting the Michaelis-Menten model in Stan [3]. A Runge-Kutta 4th/5th order integrator is used to solve the model's ODE and evaluate the likelihood. While some chains finish in seconds, others require more than an hour. Additional analysis finds that all chains are converging to the same distribution (Figure 5.).



Figure 3. Efficiency of MCMC as measured by ESS/s. Adding more chains improves the ESS. However, **waiting for the slowest chain to finish can reduce the overall ESS/s resulting in a lost in efficiency.** The maximum Efficiency is obtained after the 13 fastest chains finished, out of 30.

Waiting for the fastest chains



Figure 4. Racing-Chains MCMC. We sample chains in parallel. Each time a chain finishes, we realize a mixture sampling draw via chain stacking [4]. We assess the ESS of this draw and check its reliability through a Pareto-k diagnostic [5]. Finally, if the draw is satisfying, we keep it and drop the other chains.

	(a) Standard MCMC	(b) Standard MCMC	(c) Racing- chains	(d) Racing-chains + stacking
Chains	30	9	30 (9 kept)	9
Running time (s)	4918	775	145	145
Effective Sample Size	2728	1172	966	1107
Efficiency (ESS/s)	0.55	1.51	6.66	7.63

Table 1. Comparison of sampling methods: (a) 30 chains in parallel. (b) Another draw launching 9 chains in parallel. (c) Keep only 9 fastest chains out of 30. (d) Keep only the fastest chains for which their stacking verifies ESS > ESS_{target} = 1000. Keeping only the fastest chains reduces computational time and increase the method's efficiency.

Chain stacking [4] - principle

- The idea of chain stacking is to use parallel chains to explore as many modes as possible and then construct a stacked inference. The weights between the chains are found by maximizing the leave-one-out log predictive density (lpd_{loo}).
- From the sampling of the chains, we can build a mixture draw according to their respective weights and estimate its Effective Sample Size. Its convergence to posterior distribution is monitored by the stability of Ipd_{Ioo} seen as a function of the number of chains in parallel.
- In our case, stacking is necessary as some chains may finish quickly but have low likelihood. A uniform draw will thus induce a bias. This is especially the case for very wide priors, and the sampler can not draw a Markov chain.



Figure 5. Posterior distribution of the standard MCMC run for 30 chains in parallel and the racing-chains MCMC coming from the chain stacking of the fastest chains. The parameters drawn are those of the Michaelis-Menten, and "lp_" describes the unnormalized log posterior of the observations.

Discussion

- An open question is to understand what bias does dropping chains introduce? See the discussion by Rosenthal (2000) [6] on failing chains.
- In our example, slow chains get stuck in pathological regions of the parameter space with no probability mass and dropping these chains doesn't incur a bias (Figure 5); for another example, see [7]. Still, we can't expect this to always be true.
- Developing reliable convergence diagnostics and using unbiased Monte Carlo methods [8, 9, 10] may provide a solution.

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